## Lecture 05 Geometry of Least Squares

16 September 2015

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## Goals for today

- 1. Geometry of least squares
- 2. Projection matrix P and annihilator matrix M
- 3. Multivariate Galton Heights

## Geometry of Least Squares

Last time, we established that the least squares solution to the model:

$$y = X\beta + \epsilon$$

Yields the solution:

$$\widehat{\beta} = (X^t X)^{-1} X^t y$$

As long as the matrix  $X^t X$  is invertable.

Define the column space of the matrix *X* as:

$$\mathcal{R}(X) = \{\theta : \theta = Xb, \ b \in \mathbb{R}^p\} \subset \mathbb{R}^n$$

This is the space spanned by the p columns of X sitting in n-dimensional space.

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Notice that the least squares problem can be re-written as:

$$\widehat{\theta} = \operatorname*{arg\,min}_{\theta} \left\{ ||y - \theta||_2^2, \quad \mathrm{s.t} \quad \theta \in \mathcal{R}(X) \right\}$$

Where then  $\widehat{\beta} = X\widehat{\theta}$ .

**Theorem 3.2 (p.g. 37, Rao & Toutenburg)** The minimum,  $\hat{\theta}$  is attained when  $(y - \hat{\theta}) \perp \mathcal{R}(X)$ . In other words,  $(y - \hat{\theta})$  is perpendicular to all vectors in  $\mathcal{R}$ .

*Proof*: Pick a  $\widehat{\theta}$  in  $\mathcal{R}$  such that  $(y - \widehat{\theta}) \perp \mathcal{R}(X)$ .

$$||y - \theta||_2^2 = (y - \widehat{\theta} + \widehat{\theta} - \theta)^t (y - \widehat{\theta} + \widehat{\theta} - \theta)$$

$$\begin{aligned} ||y - \theta||_2^2 &= (y - \widehat{\theta} + \widehat{\theta} - \theta)^t (y - \widehat{\theta} + \widehat{\theta} - \theta) \\ &= (y - \widehat{\theta})^t (y - \widehat{\theta}) + (\widehat{\theta} - \theta)^t (\widehat{\theta} - \theta) + 2(y - \widehat{\theta})^t (\widehat{\theta} - \theta) \end{aligned}$$

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So, if such a  $\hat{\theta}$  exists it attains the minimum. To see that it does, write  $\hat{\theta} = X\hat{\beta}$ .

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So, if such a  $\hat{\theta}$  exists it attains the minimum. To see that it does, write  $\hat{\theta} = X\hat{\beta}$ . Then:

$$\begin{aligned} X^t(y - \widehat{\theta}) &= X^t(y - X\widehat{\beta}) \\ &= X^t y - X^t X \widehat{\beta} \end{aligned}$$

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=  $X^{t}y - X^{t}X\widehat{\beta}$   
=  $X^{t}y - X^{t}X(X^{t}X)^{-1}X^{t}y$   
=  $X^{t}y - X^{t}y$   
=  $0$ 

And therefore our proposed  $\widehat{\theta} \in \mathcal{R}(X)$ .

From this geometric interpretation of the least squares estimator, we introduce an important matrix  $P_X$  called the *projection matrix*.

$$P_X = X(X^t X)^{-1} X^t$$

I'll often drop the subscript as it should be understood that the projection is on the data matrix X.

Notice that PX = X:

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And *Py* gives the fitted values  $\hat{y}$ :

$$Py = X(X^{t}X)^{-1}X^{t}Xy$$
$$= X\widehat{\beta}$$
$$= \widehat{\theta}$$
$$= \widehat{y}$$

Do you see why the projection matrix is called the projection matrix?

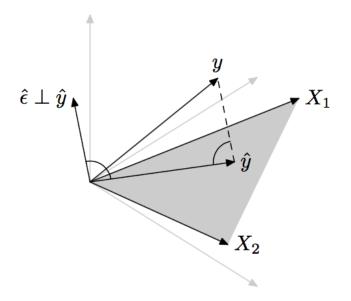
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The projection matrix is sometimes called the *hat matrix*. Any thoughts as to why?

A closely related matrix to *P* is the *annihilator matrix M*:

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It gets its name because MX = 0.

$$P^2 = X(X^t X)^{-1} X^t X(X^t X)^{-1} X^t$$

$$P^{2} = X(X^{t}X)^{-1}X^{t}X(X^{t}X)^{-1}X^{t}$$
  
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=  $X(X^{t}X)^{-1}X^{t}$   
=  $P$ 

$$M^{t} = (I_{n} - P)^{t}$$
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$$= M$$

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And idempotent:

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$$M^{t} = (I_{n} - P)^{t}$$
$$= (I_{n} - P^{t})$$
$$= M$$

And idempotent:

$$M^{2} = (I_{n} - P)^{2}$$
  
=  $(I_{n} - P)(I_{n} - P)$   
=  $I_{n} - 2 * P + P^{2}$ 

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M is also symmetric

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$$= I_{n} - 2 * P + P$$

$$= I_{n} - P$$

$$= M$$

These properties both make sense given the geometric interpretation of P and M as projections; into the column space of X and the compliment of the columns space of X.

$$r = y - X\hat{\beta}$$

$$r = y - X\widehat{\beta}$$
$$= y - Py$$

$$r = y - X\widehat{\beta}$$
  
=  $y - Py$   
=  $(I_n - P)y$ 

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$$r = y - X\widehat{\beta}$$
  
=  $y - Py$   
=  $(I_n - P)y$   
=  $My$   
=  $M(X\beta + \epsilon)$ 

$$r = y - X\widehat{\beta}$$
  

$$= y - Py$$
  

$$= (I_n - P)y$$
  

$$= My$$
  

$$= M(X\beta + \epsilon)$$
  

$$= M\epsilon$$

$$r = y - X\widehat{\beta}$$
  
=  $y - Py$   
=  $(I_n - P)y$   
=  $My$   
=  $M(X\beta + \epsilon)$   
=  $M\epsilon$ 

The matricies P and M not only help make the derivation easier, they also give geometric insight into what we are doing.

One particularly useful formula will be writing the squared residuals as:

$$||r||_{2}^{2} = ||M\epsilon||_{2}^{2}$$
$$= \epsilon^{t} M^{t} M \epsilon$$
$$= \epsilon^{t} M \epsilon$$

One particularly useful formula will be writing the squared residuals as:

$$|\mathbf{r}||_{2}^{2} = ||M\epsilon||_{2}^{2}$$
$$= \epsilon^{t} M^{t} M \epsilon$$
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So the matrix M translates the sum of squared residuals into the sum of the square errors, which are estimated by the residuals.

## **Applications**