Lecture 05
Additive Models

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Taylor B. Arnold
Yale Statistics
STAT 365/665
- Problem set notes:
  - Problem set 1 is due on Friday at 1pm!
  - brute force okay for implementation question
  - can use other libraries in the prediction and data analysis questions
  - consider FNN (R) or or sklearn.neighbors (Python)

- Office hours:
  - Taylor Arnold – Mondays, 13:00 - 14:15, HH 24, Office 206 (by appointment)
  - Elena Khusainova – Tuesdays, 13:00-15:00, HH 24, Basement
  - Yu Lu – Tuesdays, 19:00-20:30, HH 24, Basement
  - Jason Klusowski – Thursdays, 19:00-20:30, HH 24

- If you have any questions on the problem set, please ask or send them prior to Thursday night
## Factors

\[
\begin{pmatrix}
\text{Canada} \\
\text{USA} \\
\text{USA} \\
\text{Mexico} \\
\text{Canada} \\
\text{Canada} \\
\text{USA} \\
: \\
\text{Mexico}
\end{pmatrix}
\xrightarrow{}
\begin{pmatrix}
\text{Canada} & \text{Mexico} & \text{USA} \\
1 & 0 & 0 \\
0 & 0 & 1 \\
0 & 0 & 1 \\
0 & 1 & 0 \\
1 & 0 & 0 \\
1 & 0 & 0 \\
: & : & : \\
0 & 1 & 0
\end{pmatrix}
\]
Higher dimensional problems

So far, we have only considered non-parametric estimators where the predictor variable $x_i$ is one dimensional. How can we extend this to higher dimensional models?
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Well, the knn and kernel smoother estimators only depend on the distance matrix between points. Our efficient computational methods breakdown in higher dimensions, but the theoretical idea of these need no modification.
FIGURE 2.2. The same classification example in two dimensions as in Figure 2.1. The classes are coded as a binary variable (\textcolor{blue}{\textbf{BLUE}} = 0, \textcolor{orange}{\textbf{ORANGE}} = 1) and then fit by 15-nearest-neighbor averaging as in (2.8). The predicted class is hence chosen by majority vote amongst the 15-nearest neighbors.
FIGURE 2.3. The same classification example in two dimensions as in Figure 2.1. The classes are coded as a binary variable (BLUE = 0, ORANGE = 1), and then predicted by 1-nearest-neighbor classification.
Higher dimensional problems, cont.

What happens if we try to do basis expansion for linear regression in higher dimensions?
Higher dimensional problems, cont.

What happens if we try to do basis expansion for linear regression in higher dimensions?

For clarity, let’s assume we just have two dimensions labeled $x$ and $z$. We need a basis that looks like:

$$y_i = \sum_{j=0}^{m} \sum_{k=0}^{m} \beta_{j+m \cdot k} x_i^j z_i^k + \epsilon_i$$

So the number of coordinates grows to $m^2$ coefficients in order to fit arbitrary $m$ dimensional polynomials.
Higher dimensional problems, cont.

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So the number of coordinates grows to $m^2$ coefficients in order to fit arbitrary $m$ dimensional polynomials.

For $p$ dimensions, we’ll need a total $m^p$ coefficients; a quickly unfeasible task. Regularization and keeping $m$ small can help, but still makes this task hard for anything that would approximate a reasonably complex non-linear surface.
Higher dimensional problems, cont.

Lowess, local polynomial regression, can be fit in the same manor as the linear model in higher dimensions. We can fix the order of the polynomial to be $m = 1$ while still capturing global non-linearity; therefore we can still use this technique in higher dimensions.
Additive models

One way to deal with the problem of basis expansion in higher dimensions is to assume that there are no interaction between the variables. This leads to a model such as:

\[ y_i = g_1(x_{i,1}) + g_2(x_{i,2}) + \cdots + g_p(x_{i,p}) + \epsilon_i \]

These are known as additive models.
Additive models, cont.

Notice that the additive model cannot be defined uniquely as we can add a constant to one of the $g_j(\cdot)$ functions and subtract the same constant from another function $g_k(\cdot)$. In order to remedy this, one usually instead writes an explicit intercept term:

$$y_i = \alpha + g_1(x_{i,1}) + g_2(x_{i,2}) + \cdots + g_p(x_{i,p}) + \epsilon_i$$

And constrains:

$$\sum_k g_k(x_{i,k}) = 0$$

For all values of $k$. 
Computing Additive models

The primary algorithm used for computing additive models is called the backfitting algorithm. It was originally used for additive models by Leo Breiman and Jerome Friedman:

Computing Additive models, cont.

The algorithm can be compactly described as:

**Data:** pairs of data \( \{(X_i, y_i)\}_{i=1}^n \)

**Result:** Estimates \( \hat{\alpha} \) and \( \hat{g}_j, j = \{1, 2, \ldots, p\} \)

initialize \( \hat{\alpha} = \frac{1}{n} \sum_i y_i, \hat{g}_j = 0 \);

while not converged do

for \( j=1 \) to \( p \) do

\[ r_{ij} \leftarrow y_i - \hat{\alpha} - \sum_{k\neq j} \hat{g}_k(x_{ik}) \]

\[ \hat{g}_j \leftarrow S \left( \{(x_{ij}, r_{ij})\}_{i=1}^n \right) \]

\[ \hat{g}_j \leftarrow \hat{g}_j - \frac{1}{n} \sum_i \hat{g}_j(x_{ij}) \]

end

end

For some smoother function \( S \) and stopping criterion.
Computing Additive models, cont.

For the smoothing function $S$, we can use any of the algorithms we have already studied. Local polynomial regression is a popular choice.
Computing Additive models, cont.

For the smoothing function $S$, we can use any of the algorithms we have already studied. Local polynomial regression is a popular choice.

Notice that we can also blend the additive model with higher dimensional smoothers, particularly if we know that a small set of variables may have interactions with each other even though most variables do not:

$$y_i = \alpha + g_1(x_{i,1}, x_{i,2}) + g_3(x_{i,3}) + \cdots + g_p(x_{i,p}) + \epsilon_i.$$
Computing Additive models, cont.

There are two popular R packages for fitting additive models. Either `mgcv`:

https://cran.r-project.org/web/packages/mgcv

Or `gam`:

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Computing Additive models, cont.

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\end{verbatim}

Or \texttt{gam}:

\begin{verbatim}
https://cran.r-project.org/web/packages/gam
\end{verbatim}

There are not as many options for python. The best I know of is in \texttt{statsmodels.sandbox.gam} as \texttt{AdditiveModel}.
What’s wrong with linear regression?

At this point you may wonder why linear regression seems to have trouble in higher dimensions compared to the local methods. In truth, all of these techniques have trouble with high dimensional spaces; it is just that the others hide this fact in their definitions.
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The problem is the curse of dimensionality: When we have high dimensional spaces, datasets look sparse even when the number of samples is very large.
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At this point you may wonder why linear regression seems to have trouble in higher dimensions compared to the local methods. In truth, all of these techniques have trouble with high dimensional spaces; it is just that the others hide this fact in their definitions.

The problem is the curse of dimensionality: When we have high dimensional spaces, datasets look sparse even when the number of samples is very large.

Dealing with this is going to be the motivating problem in machine learning for the remainder of the course.
Description

Today we are going to look at housing price data, taking from the American Community Survey and prepared by Cosma Shalizi:

http://www.stat.cmu.edu/~cshalizi/uADA/13/hw/01/calif_penn_2011.csv

The data list aggregate statistics for census tracts.
Let’s first read in the data and look at all of the available variables.

```r
> x <- read.csv("data/CAPA.csv", as.is=TRUE)
> names(x) <- tolower(names(x))
> str(x)
'data.frame': 11275 obs. of 34 variables:
$ x : int 1 2 3 4 5 6 7 8 9 10 ... 
$ geo.id2 : num 6e+09 6e+09 6e+09 6e+09 6e+09 ... 
$ statefp : int 6 6 6 6 6 6 6 6 6 6 ... 
$ countyfp : int 1 1 1 1 1 1 1 1 1 1 ... 
$ tractce : int 400100 400200 400300 400400 400500 .. 
$ population : int 2937 1974 4865 3703 3517 1571 4206 3594 2302 5678 ... 
$ latitude : num 37.9 37.8 37.8 37.8 37.8 ... 
$ longitude : num -122 -122 -122 -122 -122 ... 
$ geo.display.label : chr "Census Tract 4001, Alameda County, California" ... 
$ median_house_value : int NA 909600 748700 773600 579200 439300 369800 ... 
$ total_units : int 1425 929 2655 1911 1703 781 1977 1738 1202 2665 ... 
$ vacant_units : int 162 37 134 68 71 65 236 257 80 500 ... 
$ median_rooms : num 6.5 6 4.6 5 4.5 4.8 4.3 4.3 4.4 4.6 ... 
```
$ mean_{household\_size\_owners}$ : num 2.02 2.53 2.45 2.04 2.66 2.72 2.17 2.7 2.75 ...
$ mean_{household\_size\_renters}$: num 1.59 1.81 1.66 2.19 1.72 2.18 2.15 1.93 1.92 2.08 ...
$ built\_2005\_or\_later$ : num 4.6 0 0 0 0 0 13.1 0 0 ...
$ built\_2000\_to\_2004$ : num 9.3 1.2 0 0.2 0.2 0 0.6 4.1 2.2 2.2 ...
$ built\_1990s$ : num 50.9 0 2.3 1.3 1.1 1.2 1.8 1.6 0.6 0 ...
$ built\_1980s$ : num 2.5 1.3 3.2 0 1.9 1.4 2.2 2.4 5.9 0.5 ...
$ built\_1970s$ : num 4.8 6.1 5.2 4.9 3.7 1 3.3 7.8 0 4.3 ...
$ built\_1960s$ : num 1.3 6.5 8.3 4.3 5.8 6.5 0.8 3.7 5.5 11.2 ...
$ built\_1950s$ : num 13.9 1 5.3 8 19.7 9.4 7.5 9.1 11.3 ...
$ built\_1940s$ : num 2.8 10.8 7.8 10.4 7.5 17 9.7 13.3 14.7 8.5 ...
$ built\_1939\_or\_earlier$ : num 9.9 73.2 68 71.1 73.8 53.1 72.4 46.5 62 62.1 ...
$ bedrooms\_0$ : num 3.6 3 11.5 5.2 4.9 3.5 8.2 8.9 14.2 6.1 ...
$ bedrooms\_1$ : num 5.6 16.4 28.4 27.7 30.2 20.4 22.3 25 20.1 29.3 ...
$ bedrooms\_2$ : num 11.9 27.4 29.2 33.7 38.1 40.1 43.2 37.5 39.4 35.4 ...
$ bedrooms\_3$ : num 40.6 34.4 20.4 21.9 19.3 30.7 16.7 25 18.3 25.3 ...
$ bedrooms\_4$ : num 31.6 17.5 7.9 7.3 5.4 4.6 6.5 2.1 5.5 3.9 ...
$ bedrooms\_5\_or\_more$ : num 6.7 1.2 2.7 4.2 2.1 0.8 3.1 1.4 2.5 0 ...
$ owners$ : num 81.2 66 45.1 45 43.6 51 32.2 28.3 31.7 35.1 ...
$ renters$ : num 18.8 34 54.9 55 56.4 49 67.8 71.7 68.3 64.9 ...
$ median\_household\_income$ : int 156250 111667 66094 87306 62386 55658 40402 ...
$ mean\_household\_income$ : int 237805 195229 105877 106248 74602 73933 ...
There are a few bad rows of data, but we can safely clean them out:

```r
> badRows <- (apply(is.na(x),1,sum) != 0)
> table(badRows)
badRows
FALSE   TRUE
 10605   670
> tapply(x$median_household_income, badRows, median, na.rm=TRUE)
  FALSE   TRUE
55459  112813
> tapply(x$median_house_value, badRows, median, na.rm=TRUE)
  FALSE   TRUE
311100  516500
> tapply(x$vacant_units, badRows, median, na.rm=TRUE)
  FALSE   TRUE
   107     70
> x <- na.omit(x)
```
As you may have guessed from the file name, the housing prices cover two distinct regions:

\[
\text{plot}(x$\$\text{longitude}, x$\$\text{latitude}, pch=19, cex=0.5)
\]
Let’s split these two states up into two separate datasets. I’ll use the California set to
start, but hopefully we will have time to go back to the Pennsylvania set.

> ca <- x[x$statefp==6,]
> pa <- x[x$statefp==42,]
As a warm-up to additive models, let’s fit and tune simple knn model for whether the majority of residents in a census tract.

```r
> testFlag <- (runif(nrow(ca)) > 0.8)
> trainFlag <- !testFlag
> cl <- as.numeric(ca$owners < 50)
```

For the training set, we will use cross-validation to select the optimal $k$:

```r
> X <- cbind(ca$latitude, ca$longitude)[trainFlag,]
> y <- cl[trainFlag]
> foldId <- sample(1:5, nrow(X), replace=TRUE)
```
Here is the main validation code, using misclassification error:

```r
> kvals <- 1:25
> res <- matrix(ncol=5, nrow=25)
> for (i in 1:5) {
+    trainSet <- which(foldId != i)
+    validSet <- which(foldId == i)
+    for (k in 1:25) {
+        pred <- knn(X[trainSet,], X[validSet,], y[trainSet],
+                      k=kvals[k])
+        yhat <- (as.numeric(pred) - 1)
+        res[k,i] <- mean((y[validSet] != yhat))
+        print(k)
+    }
+}
```
Taking the results for each fold, I can calculate the cross validated mis-classification rate as well as the standard errors of these rates:

```r
> head(res)
[1,] 0.27 0.27 0.26 0.27 0.28
[2,] 0.23 0.25 0.23 0.25 0.27
[3,] 0.24 0.25 0.26 0.25 0.27
[4,] 0.22 0.23 0.24 0.25 0.25
[5,] 0.23 0.24 0.26 0.25 0.26
[6,] 0.23 0.24 0.24 0.26 0.25
> cvError <- apply(res,1,mean)
> cvSe <- apply(res,1,sd) / sqrt(5)
```
If we set the tuning parameter to 4, we can then check how well this performs on the test set.

```r
> Xtest <- cbind(ca$latitude, ca$longitude)[testFlag,]
> ytest <- cl[testFlag]
> yhat <- (as.numeric(knn(X, Xtest, y, k=4)) - 1)
> mean((yhat != ytest))
[1] 0.22
> round(table(yhat, ytest) / length(yhat) * 100)
   ytest
yhat    0  1
  0 56 16
  1  6 22
```

The table at the bottom is called a **confusion matrix**, and gives more granularity than the raw misclassification rate.
If we set the tuning parameter to 4, we can then check how well this performs on the test set.

```r
> Xtest <- cbind(ca$latitude, ca$longitude)[testFlag,]
> ytest <- cl[testFlag]
> yhat <- (as.numeric(knn(X, Xtest, y, k=4)) - 1)
> mean((yhat != ytest))
[1] 0.22
> round(table(yhat, ytest) / length(yhat) * 100)
    yhat  ytest
   0    56  16
   1     6  22
```

The table at the bottom is called a **confusion matrix**, and gives more granularity than the raw misclassification rate.
Now, I want to understand the variables that effect the median house value in a census tract. Here is a linear model that would be a good starting point (after some exploratory plots, preferably):

```r
> ca.lm <- lm(log(median_house_value) ~ median_household_income
+   + mean_household_income + population + total_units +
+   + vacant_units + owners + median_rooms +
+   + mean_household_size_owners + mean_household_size_renters
+   + latitude + longitude, data = ca, subset=trainFlag)
```
> summary(ca.lm)

| Estimate | Std. Error | t value | Pr(>|t|) |
|----------|------------|---------|---------|
| (Intercept) | -5.78e+00 | 5.93e-01 | -9.74 | < 2e-16 *** |
| median_household_income | 1.20e-06 | 5.19e-07 | 2.30 | 0.021 * |
| mean_household_income | 1.08e-05 | 4.35e-07 | 24.73 | < 2e-16 *** |
| population | -4.15e-05 | 5.59e-06 | -7.42 | 1.3e-13 *** |
| total_units | 8.37e-05 | 1.73e-05 | 4.83 | 1.4e-06 *** |
| vacant_units | -1.06e-06 | 2.64e-05 | -0.04 | 0.968 |
| owners | -3.83e-03 | 3.57e-04 | -10.72 | < 2e-16 *** |
| median_rooms | -1.49e-02 | 9.36e-03 | 1.59 | 0.112 |
| mean_household_size_owners | 5.40e-02 | 7.99e-03 | 6.76 | 1.5e-11 *** |
| mean_household_size_renters | -7.46e-02 | 7.20e-03 | -10.36 | < 2e-16 *** |
| latitude | -2.15e-01 | 6.36e-03 | -33.81 | < 2e-16 *** |
| longitude | -2.15e-01 | 6.67e-03 | -32.29 | < 2e-16 *** |

---

Signif. codes: 0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ‘ 1

Residual standard error: 0.32 on 5995 degrees of freedom
Multiple R-squared:  0.636, Adjusted R-squared:  0.635
F-statistic: 953 on 11 and 5995 DF,  p-value: <2e-16
To fit an additive model in R, we can use the `mgcv` package. It uses cross-validation by default, making it very easy to use in place of linear regression.

```r
> library(mgcv)
Loading required package: nlme
This is mgcv 1.8-7. For overview type 'help("mgcv-package")'.
> ca.gam <- gam(log(median_house_value)
+ ~ s(median_household_income) + s(mean_household_income)
+ + s(population) + s(total_units) + s(vacant_units)
+ + s(owners) + s(median_rooms) + s(mean_household_size_owners)
+ + s(mean_household_size_renters) + s(latitude)
+ + s(longitude), data=ca, subset=trainFlag)
```
To see the ‘coefficients’ in the additive model, we can plot the output object. These options work well when working locally:

```r
> plot(ca.gam2, scale=0, se=2, shade=TRUE, resid=FALSE, pages=1)
```

For class, I add the option `select=i` to only show the contribution of the $i$th variable.
 owners
s(owners, 3.87)

 owners
It actually makes more sense to allow and interaction between latitude and longitude. This is also easy to include in mgcv:

```r
> ca.gam2 <- gam(log(median_house_value)
+ ~ s(median_household_income) + s(mean_household_income)
+ + s(population) + s(total_units) + s(vacant_units)
+ + s(owners) + s(median_rooms) + s(mean_household_size_owners)
+ + s(mean_household_size_renters)
+ + s(longitude,latitude), data=ca, subset=trainFlag)
```
How well does these methods do in terms of prediction? We can predict using the `predict` function just as with linear models:

```r
> y <- log(ca$median_house_value)
> ca.lm.pred <- predict(ca.lm, ca)
> ca.gam.pred <- predict(ca.gam, ca)
> ca.gam2.pred <- predict(ca.gam2, ca)
```

And then check the mean squared error on both the training set and testing set:

```r
> tapply((ca.lm.pred - y)^2, trainFlag, mean)
FALSE   TRUE
0.096   0.101
> tapply((ca.gam.pred - y)^2, trainFlag, mean)
FALSE   TRUE
0.064   0.072
> tapply((ca.gam2.pred - y)^2, trainFlag, mean)
FALSE   TRUE
0.059   0.065
```
In machine learning, you’ll often hear the caveat that everything depend on future values following the same underlying model. I think we say that a lot, but forget to really think about it. To illustrate, let’s re-fit the model on the California data without the latitude and longitude components. We can then see how well the model trained on California data generalizes to Pennsylvania data.
Here are the two linear models fit on the two different datasets.

```r
> ca.lm2 <- lm(log(median_house_value) ~ median_household_income
+     + mean_household_income + population + total_units +
+     + vacant_units + owners + median_rooms +
+     + mean_household_size_owners + mean_household_size_renters,
+     + data = ca, subset=trainFlag)
>
> pa.lm3 <- lm(log(median_house_value) ~ median_household_income
+     + mean_household_income + population + total_units +
+     + vacant_units + owners + median_rooms +
+     + mean_household_size_owners + mean_household_size_renters,
+     + data = pa, subset=trainFlag)
```
And here are the two additive models fit on the data:

```r
> ca.gam3 <- gam(log(median_house_value) + s(median_household_income) + s(mean_household_income) + s(population) + s(total_units) + s(vacant_units) + s(owners) + s(median_rooms) + s(mean_household_size_owners) + s(mean_household_size_renters), data=ca, subset=trainFlagPa)
> pa.gam4 <- gam(log(median_house_value) + s(median_household_income) + s(mean_household_income) + s(population) + s(total_units) + s(vacant_units) + s(owners) + s(median_rooms) + s(mean_household_size_owners) + s(mean_household_size_renters), data=pa, subset=trainFlagPa)
```
Fitting these models all on the PA data:

```r
> y.pa <- log(pa$median_house_value)
> pa.lm2.pred <- predict(ca.lm2, pa)
> pa.gam3.pred <- predict(ca.gam3, pa)
> pa.lm3.pred <- predict(pa.lm3, pa)
> pa.gam4.pred <- predict(pa.gam4, pa)
```

We see that the California ones yield very poor MSE scores for PA:

```r
> tapply((pa.lm2.pred - y.pa)^2, trainFlagPa, mean)
FALSE  TRUE
   0.58 0.55
> tapply((pa.gam3.pred - y.pa)^2, trainFlagPa, mean)
FALSE  TRUE
   0.47 0.44
> tapply((pa.lm3.pred - y.pa)^2, trainFlagPa, mean)
FALSE  TRUE
   0.095 0.093
> tapply((pa.gam4.pred - y.pa)^2, trainFlagPa, mean)
FALSE  TRUE
   0.070 0.063
```
If we account for the overall means being different, we see that the California models perform reasonably well on the Pennsylvania data:

```r
> tapply((pa.lm2.pred - y.pa), trainFlagPa, var)
  FALSE  TRUE
0.14 0.13
> tapply((pa.gam3.pred - y.pa), trainFlagPa, var)
  FALSE  TRUE
0.094 0.084
> tapply((pa.lm3.pred - y.pa), trainFlagPa, var)
  FALSE  TRUE
0.095 0.093
> tapply((pa.gam4.pred - y.pa), trainFlagPa, var)
  FALSE  TRUE
0.070 0.063
```